

## Monetary Policy Shocks: Data or Methods?

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Identifying monetary policy is difficult!

- It is endogenous: both responds to and affects the economy

High-frequency changes in asset prices proxy exogenous variation in monetary policy since Kuttner (2001)

- isolate monetary news by comparing asset prices shortly before and after FOMC announcements
- many possible data and methods → many different monetary policy shocks available
  - Eurodollar and federal funds futures, Treasuries
  - first differences, principal component analysis, Fama-MacBeth regression

## Stating the Problem

**Problem:** the sizes/signs of monetary shock series can vary across data/methods

- correlation as low as 0.3, same sign about 1/2 of the time
- differences starkest at the ELB

How do these differences affect estimates of monetary transmission?

## What We Do

Compare data/methods of Kuttner (2001), Gertler & Karadi (2015), Nakamura & Steinsson (2018)/Gürkaynak et al. (2005), Bu et al. (2021)

1. shock construction
2. monetary policy transmission

We compare shocks that are

- **high-frequency**, for VAR or narrative comparisons see: Rudebusch (1998), Ettmeier and Kriwoluzky (2019)
- **single series**, for multiple dimensions see: Gürkaynak et al. (2005), Lewis (2023), Swanson (2021, 2023), Acosta (2023), Jarocinski (2023)
- **w/out add-ons**: Miranda-Agrippino and Ricco (2021), Bauer and Swanson (2023, 2022), Jarocinski and Karadi (2020), Nunes et al. (2023), Zhu (2023)

# What We Find

Monetary transmission:

- IRFs of LPs and VARs more similar across shock series than forecast revisions
- swapping data/methods → qualitative differences: inference not robust to construction

Bu et al. (2021) shock series:

- unpredictable by economic news
- similar in ELB and non-ELB periods
- mitigates an adversely-signed response in the specifications we study

Attribute to data on long-term rates *and* a method that extract their differential responsiveness relative to short-term rates

- high-frequency shocks pioneered when the FFR was the key policy instrument (and no ELB...); now there is a portfolio of tools

# Shock Construction

# 1. Data Only: MP1 and FF4

**MP1:** Kuttner (2001)

$$MP1_s = \begin{cases} \frac{D^s}{D^s - d^s} (ff_{s,t}^1 - ff_{s,t-\Delta t}^1) & \text{if } D^s - d^s > 7 \\ ff_{s,t}^2 - ff_{s,t-\Delta t}^2 & \text{otherwise} \end{cases}$$

**FF4:** Gertler & Karadi (2015), Jarocinski & Karadi (2021)

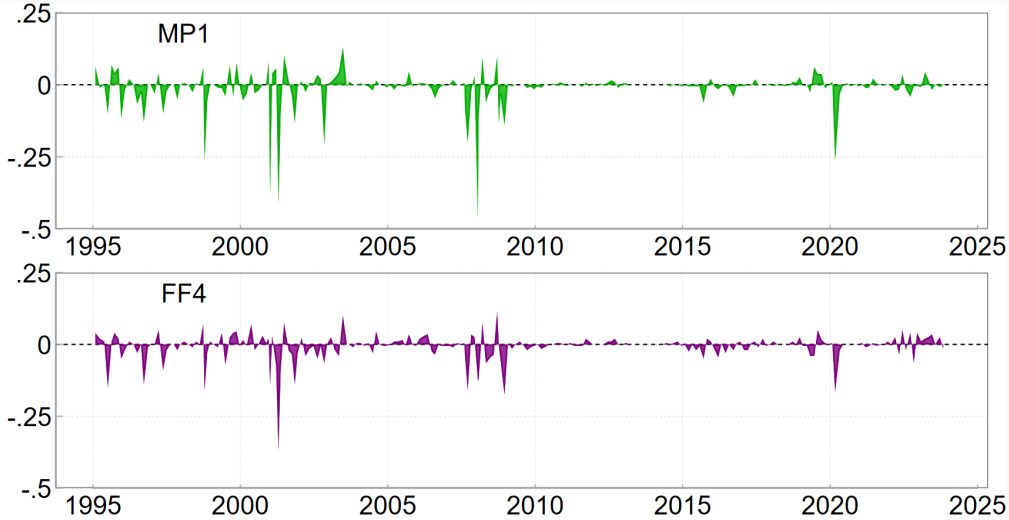
▶ next announcement

$$\Delta ff_s^4 = (ff_{s,t}^4 - ff_{s,t-\Delta t}^4)$$



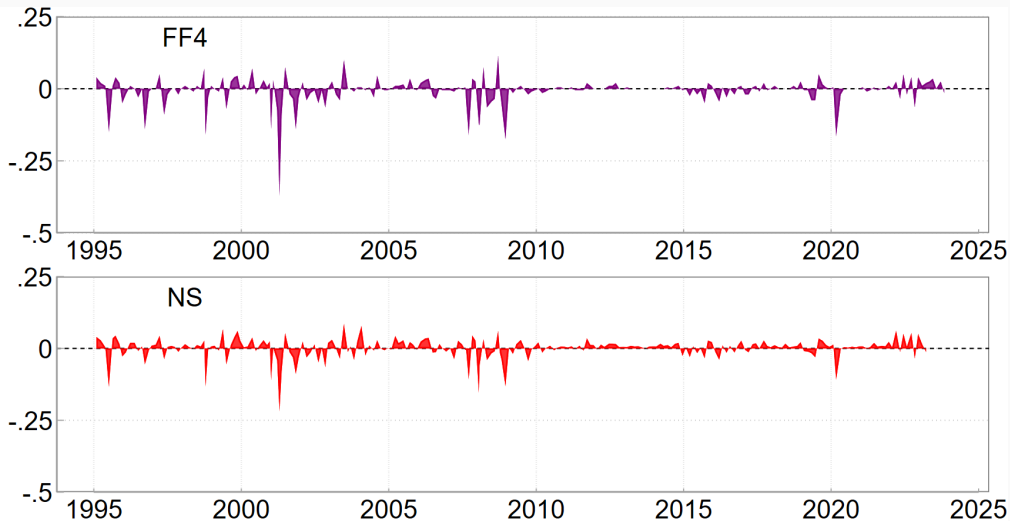
# Data Only: MP1 and FF4

▶ histograms



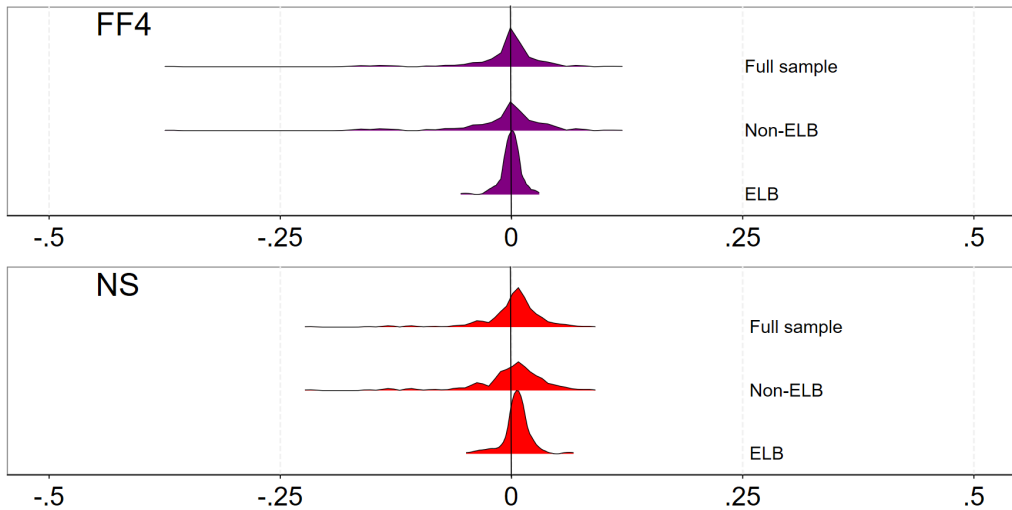
correlation=0.8 , same-sign=59%





correlation=0.92 , same-sign=66%

## Data & Methods: Nakamura and Steinsson (NS)



## Data and Methods: Bu et al. (2021)

- Fama and MacBeth (1973) regression on Treasury yield curve
- One day window surrounding FOMC announcement  $s$ 
  1. estimate  $\{\hat{\beta}_j\}_{j=1}^{30}$  via separate regressions

$$\begin{aligned}\Delta R_s^1 &= \alpha_1 + \beta_1 \Delta i_s + \epsilon_s^1 \\ &\vdots \\ \Delta R_s^{30} &= \alpha_{30} + \beta_{30} \Delta i_s + \epsilon_s^{30}\end{aligned}$$

assume  $\Delta i_s$  is one-to-one with 2-year yield  $\Delta R_s^2$ , so  $\Delta i = \Delta R_s^2 - \epsilon_s^2$ . Then,

$$\Delta R_s^j = \theta_j + \beta_j \Delta R_s^2 + \underbrace{\epsilon_s^j - \beta_j \epsilon_s^2}_{\xi_s^j}$$

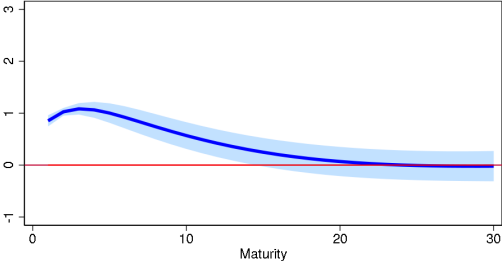
endogeneity of  $\text{corr}(\Delta R_s^j, \xi_s^j) > 0$  reconciled w/ IV or Rigobon (2003) estimator

2. estimate shocks  $\Delta \hat{i}_s$  from  $s = 1, \dots, T$  cross-sectional reg. of  $\Delta R_s^j$  on  $\{\hat{\beta}_j\}_{j=1}^{30}$

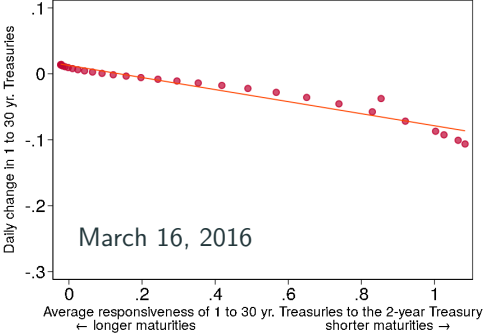
$$\Delta R_s^j = \alpha_j + \Delta \hat{i}_s \hat{\beta}_j + v_s^j, \quad s = 1, \dots, T \text{ announcements}$$

3. Re-scale shocks series to  $\Delta R_s^2$

### Average Response of Treasuries to Changes in the 2-year Treasury on FOMC Days



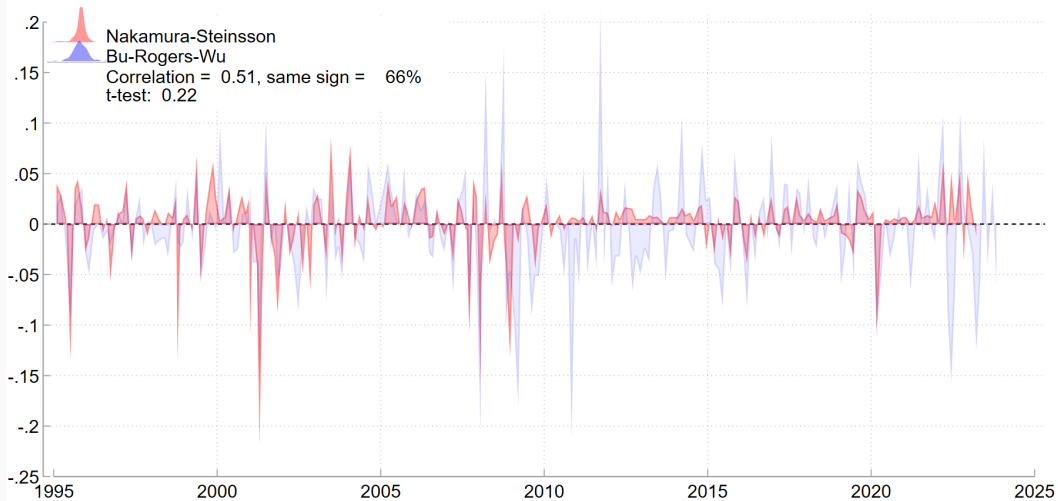
— Average Response Over Whole Sample 1995-2023  
Shaded Bands Represent 95% Confidence Intervals



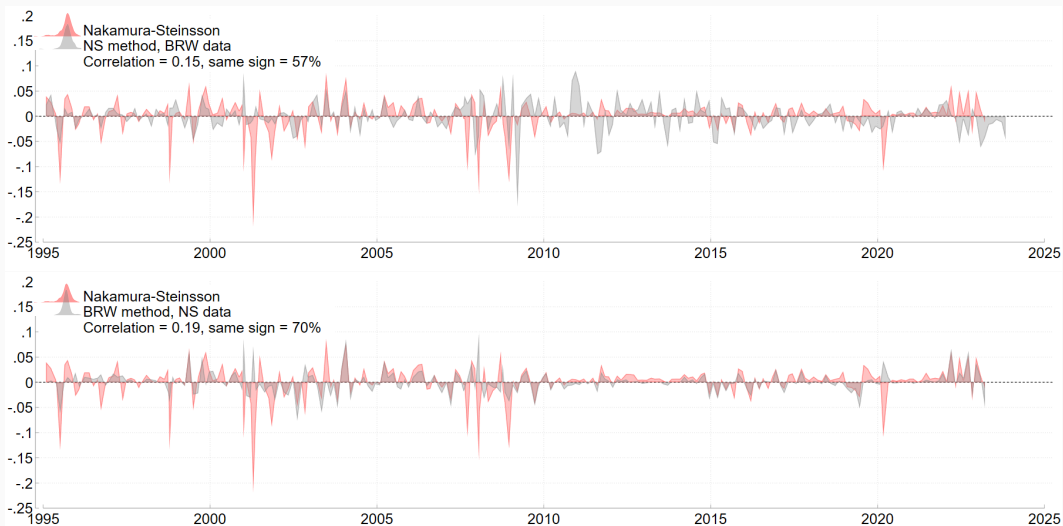
March 16, 2016

Average responsiveness of 1 to 30 yr. Treasuries to the 2-year Treasury  
← longer maturities                      shorter maturities →

# Why the difference?



### 3. Swapped shocks: differences due to methods or data?



# Estimating Monetary Transmission

Monthly revisions to Blue-chip forecasts à la Nakamura and Steinsson (2018),  
Campbell et al. (2012)

$$\text{Blue Chip GDP Revisions}_T = \beta \varepsilon_t^i + e_T$$

for  $i =$  data, data and methods, swapped

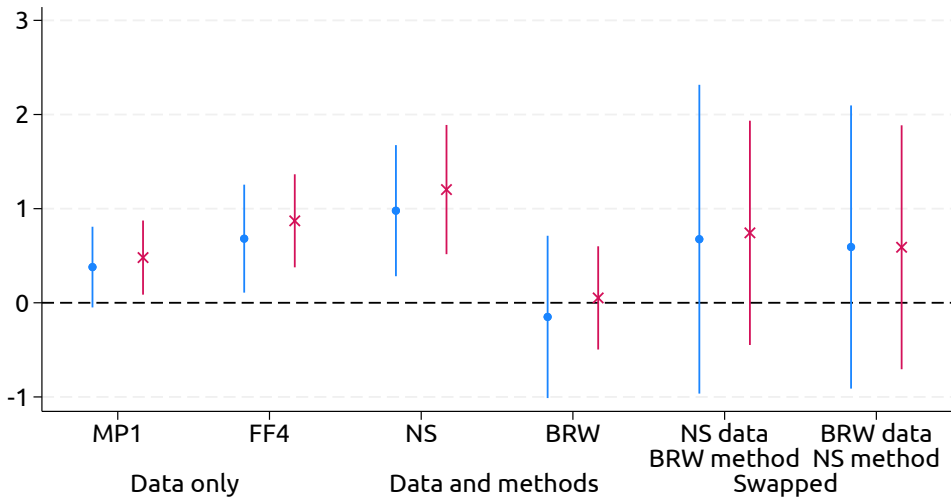
- $\hat{\beta}$  often the opposite sign of theoretical predictions



# Blue Chip Regression Coefficients, 95% CI

▶ Other samples

▶ ELB



• Full sample, 1995-2023

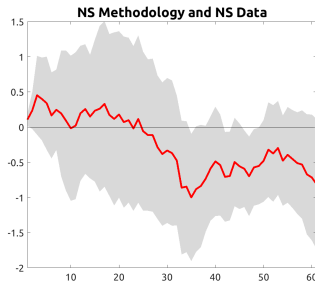
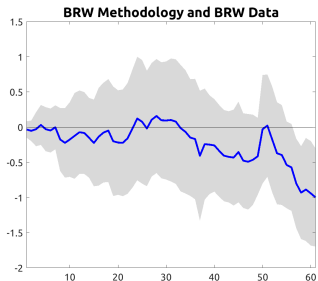
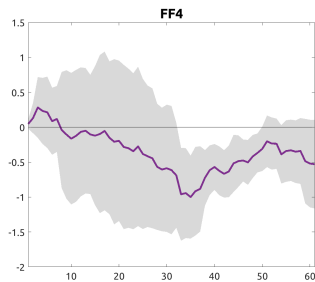
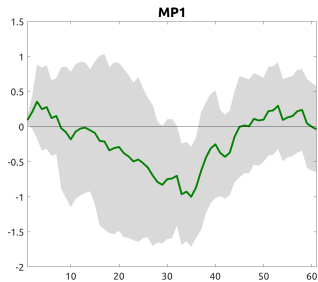
× NS sample, 1995-2016

Billion Prices Project's daily CPI in a local projection for day  $t + h$ .

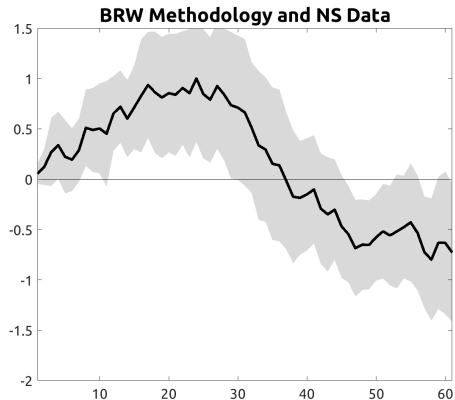
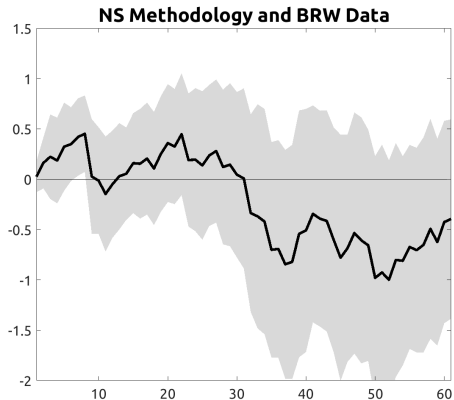
$$\pi_{t+h} = \alpha_{(h)} + \beta_{(h)} \varepsilon_t^i + \Gamma_h z_t + e_t^h$$

- LHS and RHS both high-frequency
- Jacobson et. al (2023): high-frequency LHS mitigates temporal aggregation bias

## 2. Local Projections 2008-2015, 90% CI

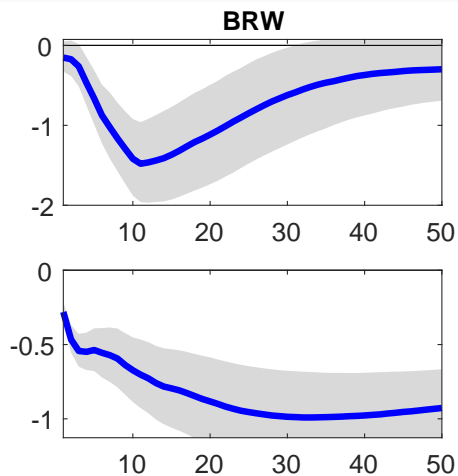
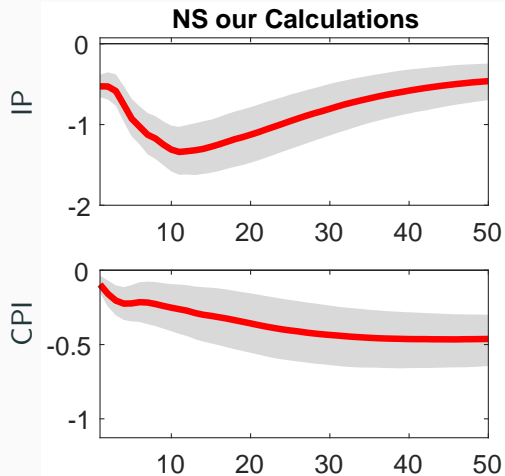


## 2. Local Projections 2008-2015, 90% CI



Gertler and Karadi (2015) VAR with the monetary shock as an external instrument à la Bauer and Swanson (2022)

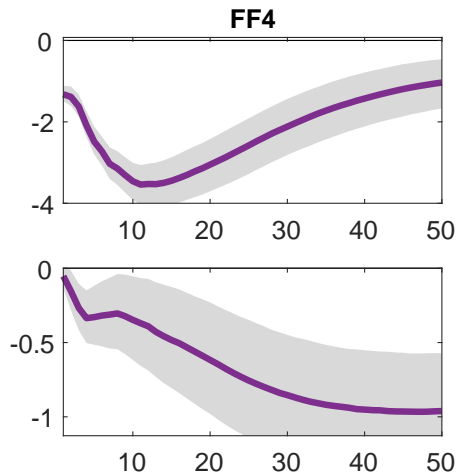
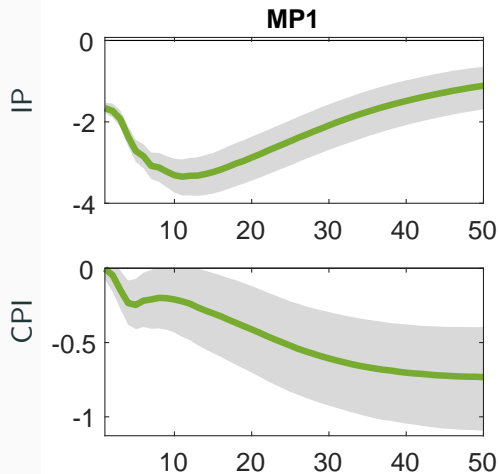
- 8 lags, 4 variables: IP, CPI, excess bond premium, 2-year Treasury
- VAR: 1973 to Feb. 2020
- External instrument: 1995 to 2019
- Canova & Ferroni Toolbox

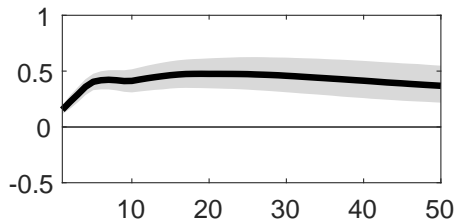
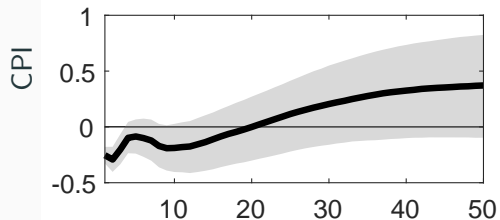
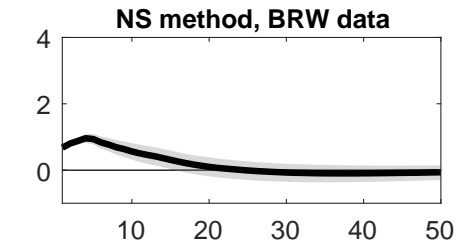
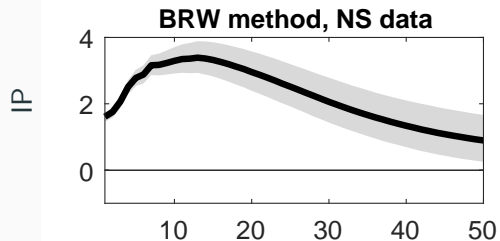


# VAR IRFs, 25 bps Monetary Shock

▸ detailed comparison

▸ financial variables







Document differences in sizes/signs of high-frequency monetary shocks

- starkest at the ELB

Attribute differences to both data and methods

- shift away from a single instrument to a portfolio of tools
- exploiting additional information from long-term rates depends on the method

Consequences for inference depends on the specification

- qualitatively similar in some LPs and VARs
- qualitatively different when swapping data/methods, inference may not be robust to all constructions

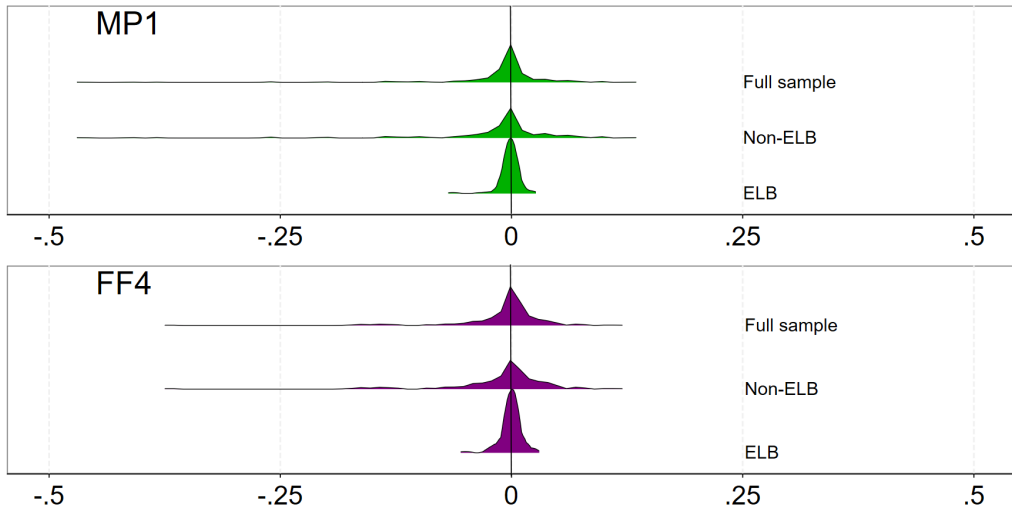
# Appendix

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Future		Percent	Number
<i>FF1</i>	in current month	2.05%	5
<i>FF2</i>	1-month ahead	50.41%	123
<i>FF3</i>	2-months ahead	47.13%	115
<i>FF4</i>	3-months ahead	0.41%	1
Total			244

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# 1. Data only: High Frequency Monetary Policy Shocks



- First principal component of intraday changes in five futures
- Fed funds futures liquid for  $\approx 3$  months ahead, eurodollars 4 to 12 months ahead  
eurodollars: dollar-denominated deposits at foreign banks
- 30-minute window surrounding FOMC announcement in month  $s$

$MP1_s$  as previously shown

$$MP2_s = \begin{cases} \frac{D^{s'}}{D^{s'} - d^{s'}} \left[ (\text{ff}_{s',t}^j - \text{ff}_{s',t-\Delta t}^j) - \frac{d^{s'}}{D^{s'}} MP1_s \right] & \text{if } D^{s'} - d^{s'} > 7, \quad j \approx 2, 3 \\ \text{ff}_{s',t}^{j+1} - \text{ff}_{s',t-\Delta t}^{j+1} & \text{otherwise} \end{cases}$$

[▶ details](#)

$$edk_q = edk_{q,t} - edk_{q,t-\Delta t}, \quad k = 2, 3, 4 \quad \text{quarters ahead}$$

# Construction of Nakamura and Steinsson (2018) Shocks

1. expected ffr ( $r_0$ ) for the month of the FOMC announcement, adjusted elapsed days of the month ( $d^s$ ) out of total days ( $D^s$ )

$$\underbrace{f_{t-\Delta t}^1}_{\text{current month's ffr future prior to FOMC}} = \underbrace{\frac{d^s}{D^s} r_{-1}}_{\text{current month's ffr prior to FOMC}} + \underbrace{\frac{D^s - d^s}{D^s} \mathbb{E}_{t-\Delta t} r_0}_{\text{current month's ffr future prior to FOMC}}$$

$$\underbrace{f_t^1}_{\text{current month's ffr future after FOMC}} = \underbrace{\frac{d^s}{D^s} r_{-1}}_{\text{current month's ffr prior to FOMC}} + \underbrace{\frac{D^s - d^s}{D^s} \mathbb{E}_t r_0}_{\text{current month's ffr future after FOMC}}$$

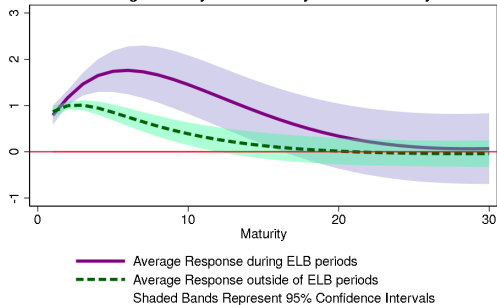
Combining and re-arranging:

$$\underbrace{\mathbb{E}_t r_0 - \mathbb{E}_{t-\Delta t} r_0}_{\text{expected } \Delta \text{ in current month's ffr}} = \frac{D^s}{D^s - d^s} (f_t^1 - f_{t-\Delta t}^1)$$

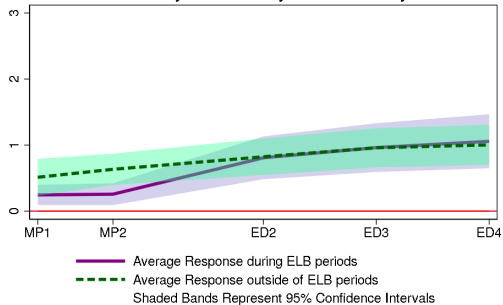
2. expected ffr  $r_1$  for the remainder of the month of the next scheduled FOMC announcement

$$\underbrace{\mathbb{E}_t r_1 - \mathbb{E}_{t-\Delta t} r_1}_{\text{expected } \Delta \text{ in ffr in month of next FOMC}} = \frac{d^{s'}}{D^{s'} - d^s} \left[ \underbrace{(f_t^n - f_{t-\Delta t}^n)}_{\Delta \text{ in ffr future for next FOMC}} - \underbrace{\frac{d^{s'}}{D^{s'}} (\mathbb{E}_t r_0 - \mathbb{E}_{t-\Delta t} r_0)}_{\text{scaled expected } \Delta \text{ in current month}} \right]$$

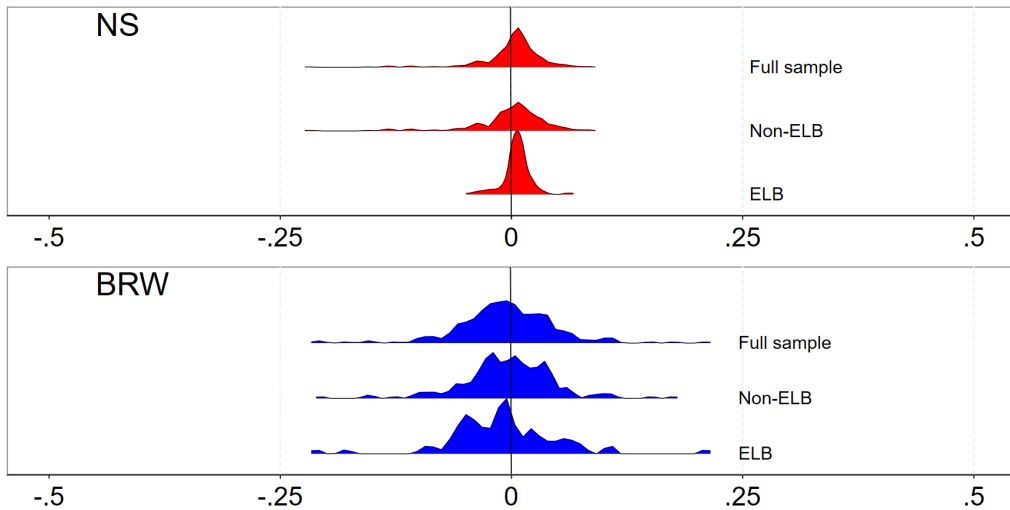
### Average Response of Treasuries to Changes in 2-year Treasury on FOMC Days



### Average Response of Short-Term Futures to Changes in the 2-year Treasury on FOMC Days

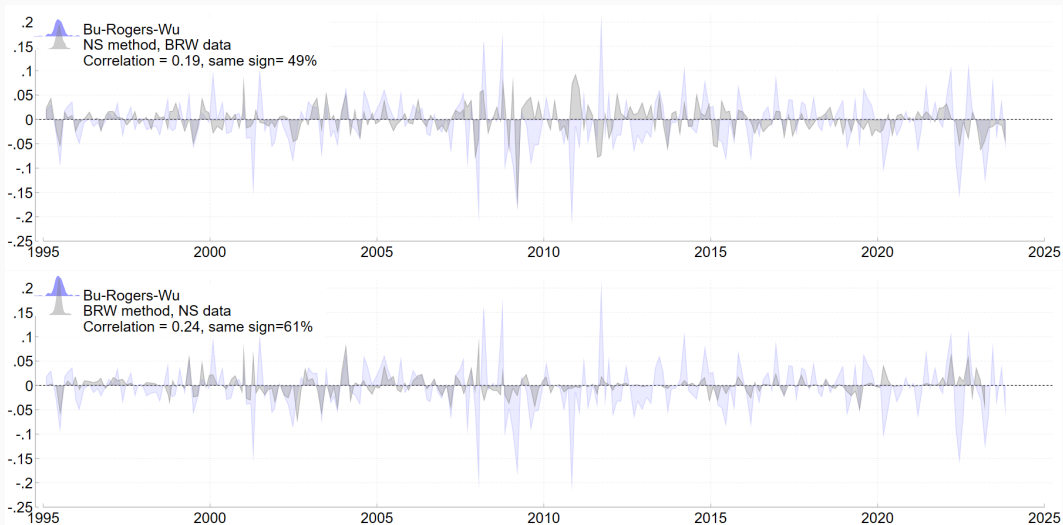


## 2. Why the difference?





### 3. Swapped shocks: differences due to methods or data?



Note: legend markers are data distributions.

Karnaukh and Vokata (2022), Sastry (2021), and Bauer and Swanson (2023) show that monetary shocks  $\varepsilon_t^i$  can be predicted by economic news

$$\varepsilon_t^i = \alpha + \beta news_t^k + e_t$$

$$news\ k = \left\{ \begin{array}{l} \text{Blue-Chip GDP revisions} \\ \text{Change in non-farm payrolls} \\ \text{Aruoba-Scotti-Diebold index} \\ \text{Brave et. al Index} \end{array} \right.$$

for  $i =$  data, data & methods, swapped shocks

# Predictability Coefficients from 1995 to 2023 (ex. Covid), 95% CI

▶ full sample

▶ full, ex crisis

▶ full, ex crisis & covid

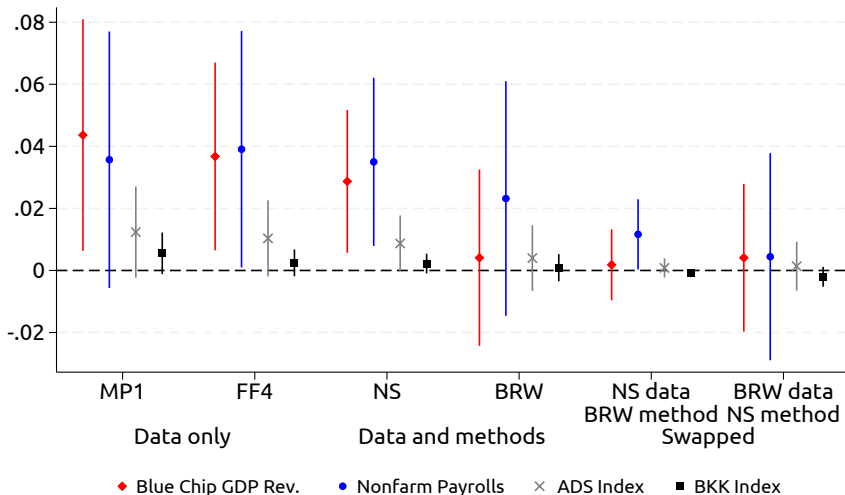
▶ 1995-2016

▶ 1995-2016, ex crisis

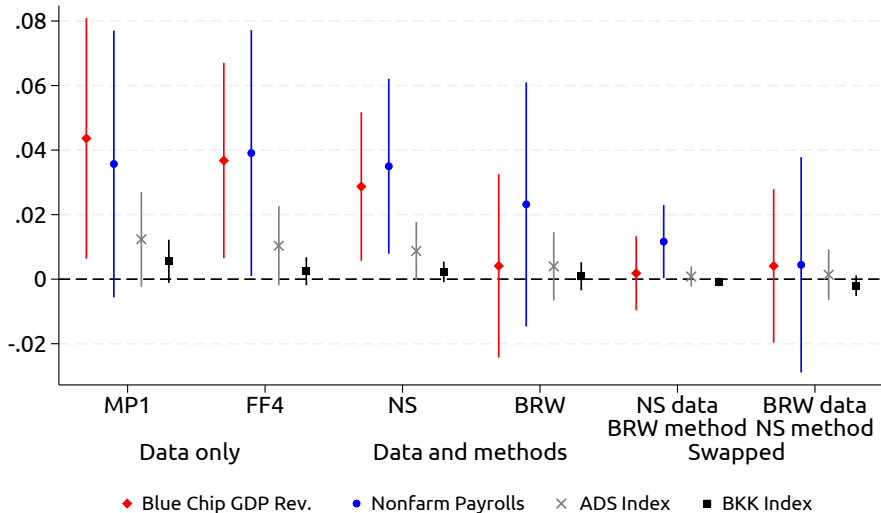
▶ ELB

▶ non-ELB

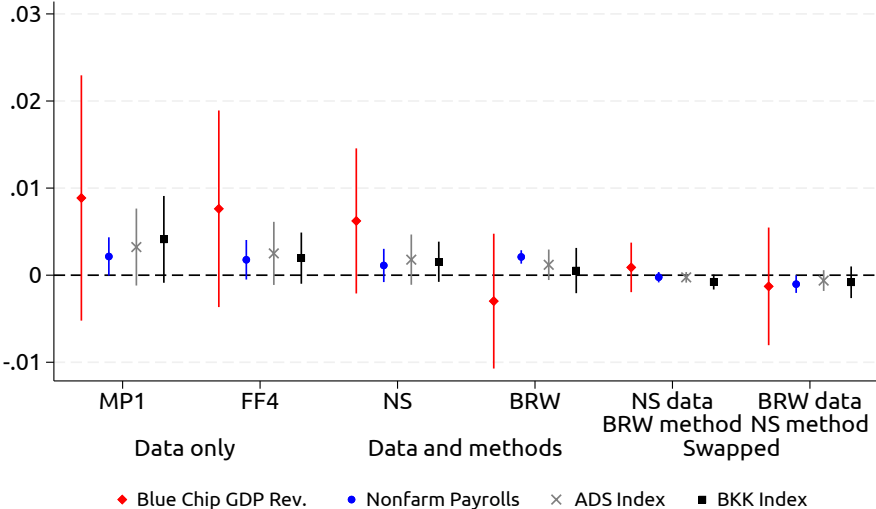
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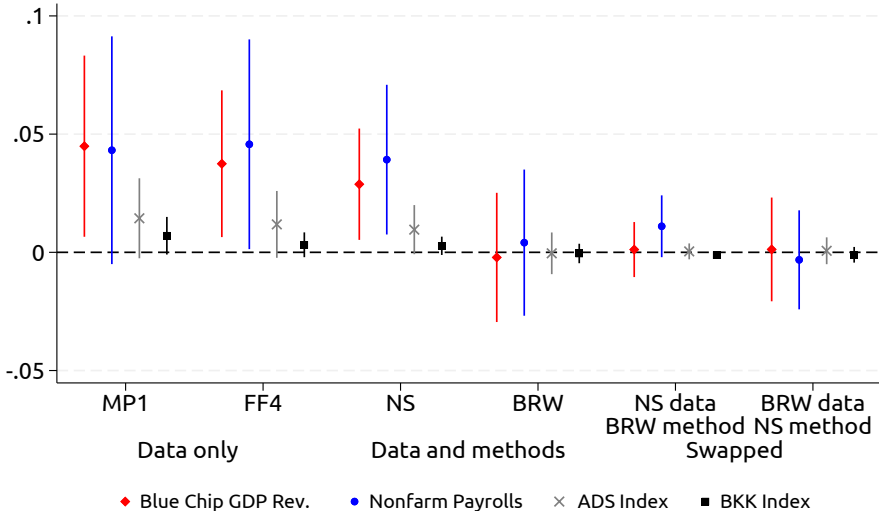
# Predictability Coefficients from 1995 to 2023, 95% CI



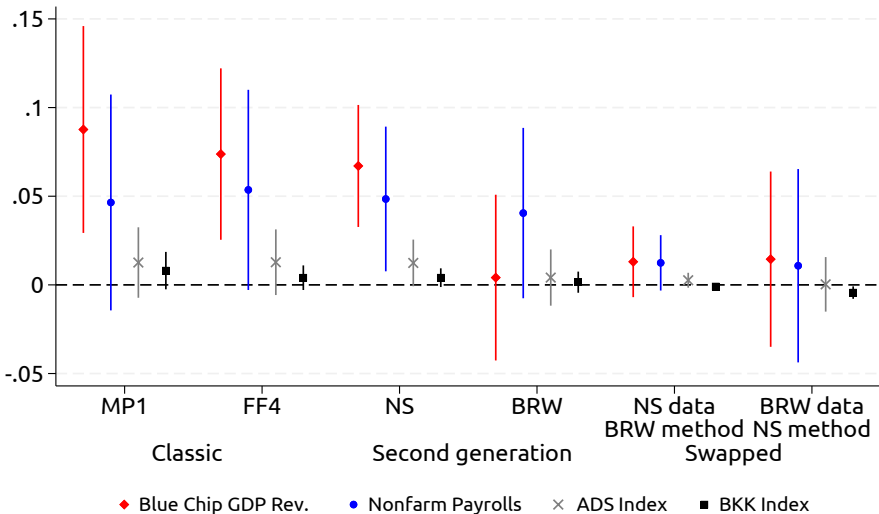
# Predictability Coefficients from 1995 to 2023 (ex. crisis), 95% CI



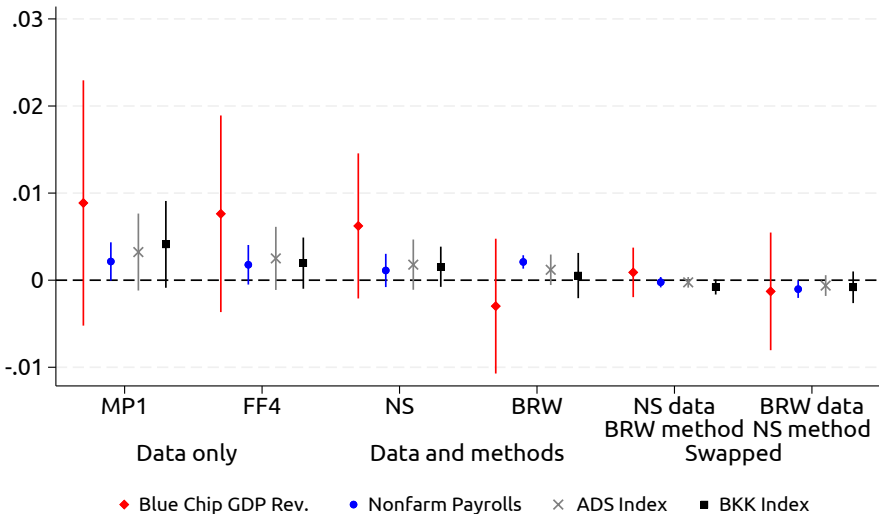
# Predictability Coefficients from 1995 to 2023 (ex. crisis & Covid), 95% CI



# Predictability Coefficients from 1995 to 2016, 95% CI

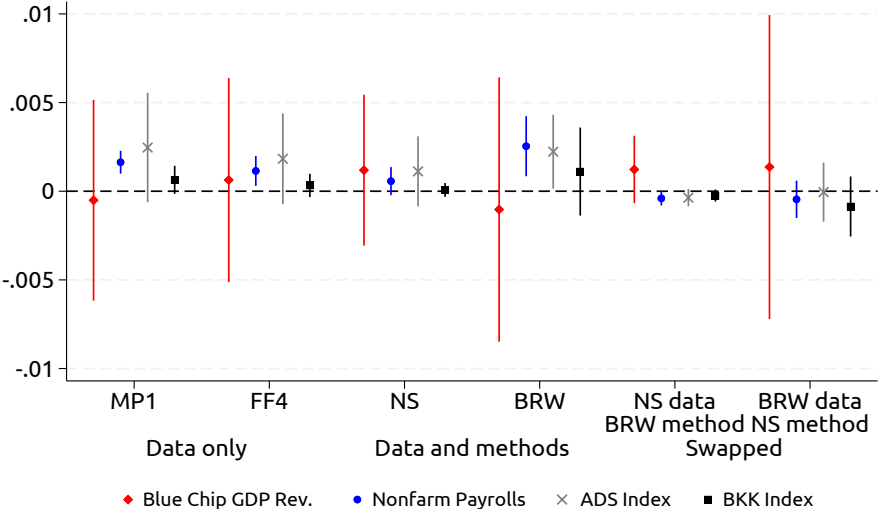


# Predictability Coefficients from 1995 to 2016 (ex. crisis), 95% CI

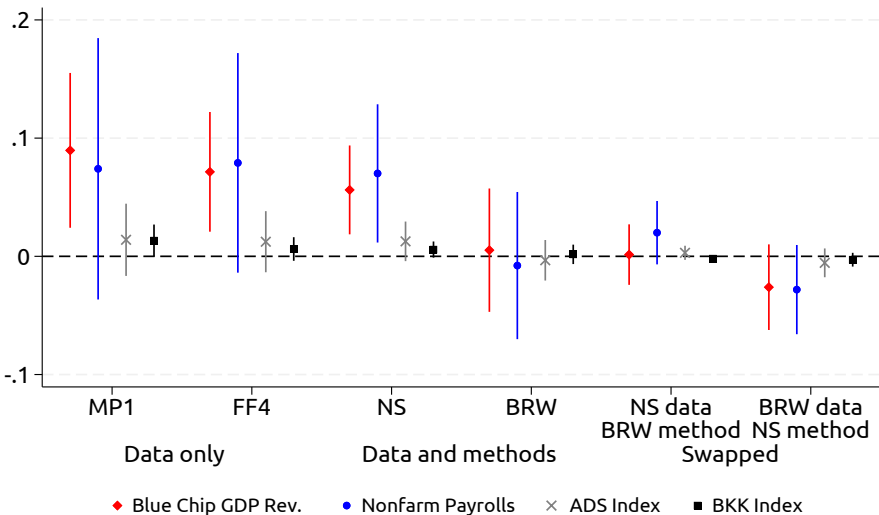




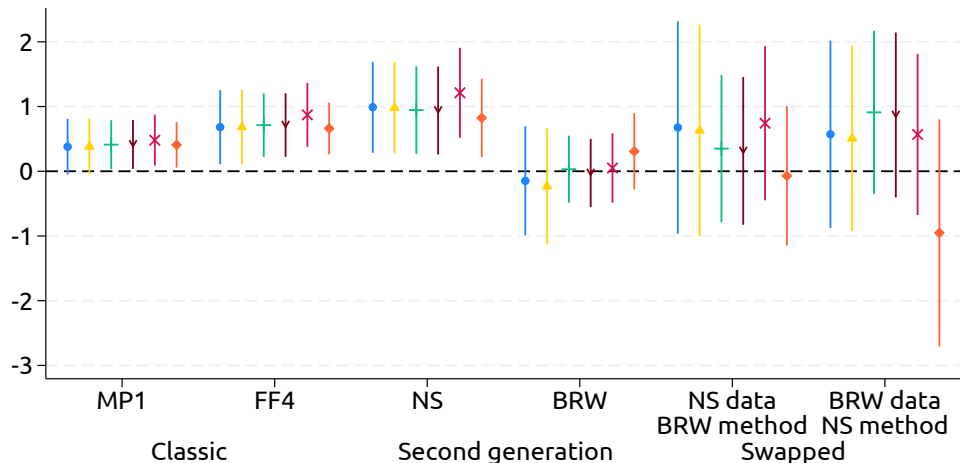
# Predictability Coefficients ELB, 95% CI



# Predictability Coefficients non-ELB, 95% CI

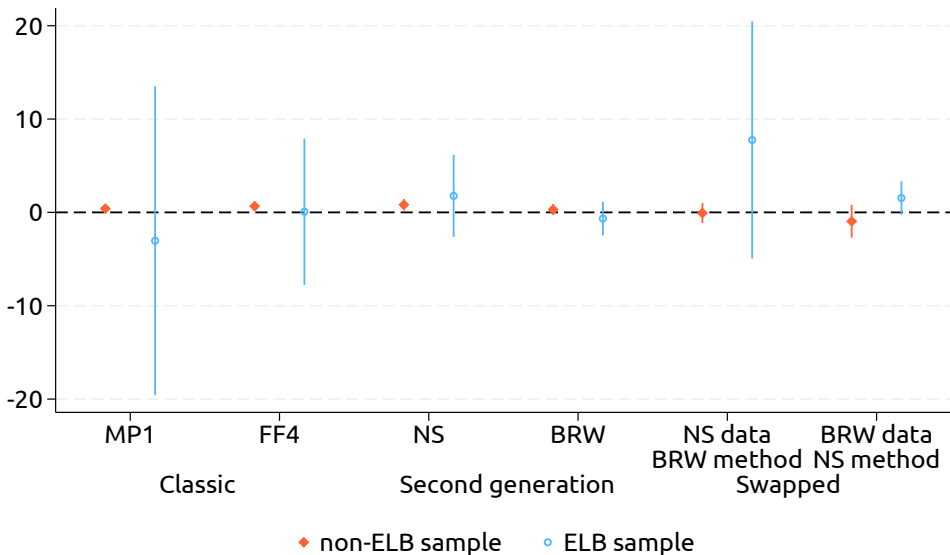


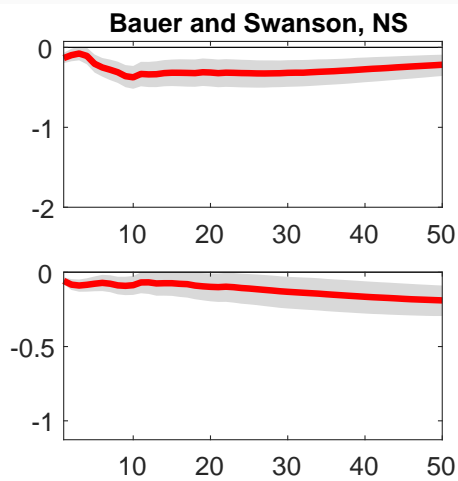
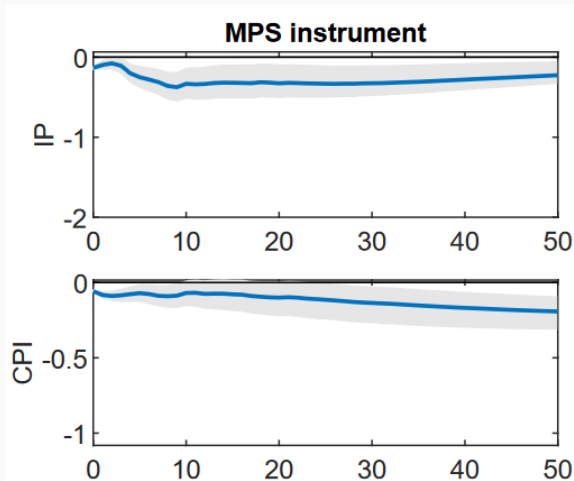
# 1. Blue Chip Regression Coefficients, 95% CI

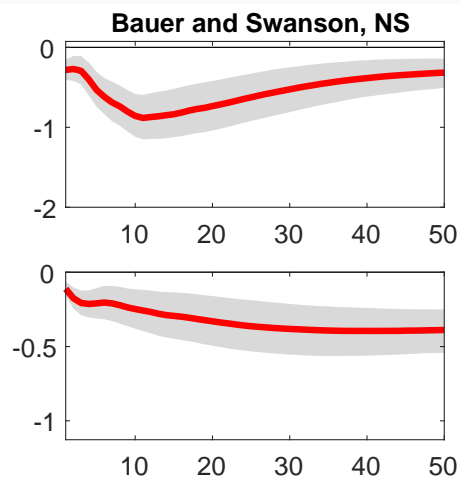
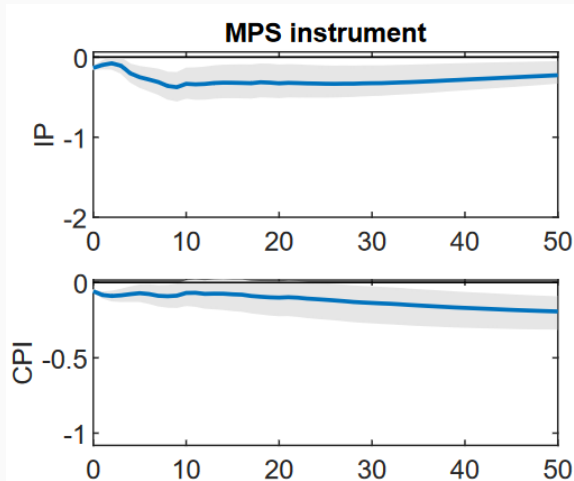


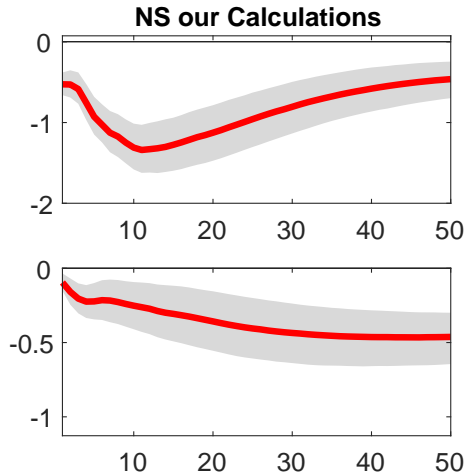
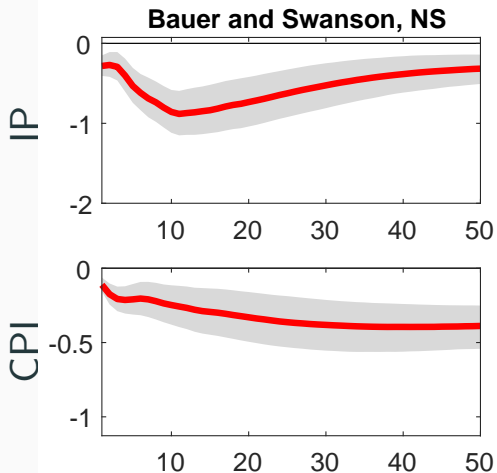
- Full, 1995-2023
- ▲ Full, crisis controls
- + Full, Covid controls
- ▼ Full, crisis and Covid controls
- \* NS Sample, 1995-2016
- ◆ non-ELB Sample

# 1. Blue Chip Regression Coefficients, 95% CI

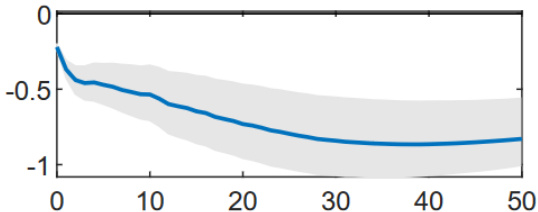
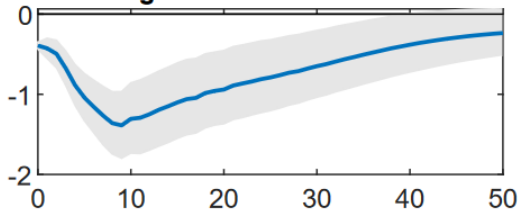




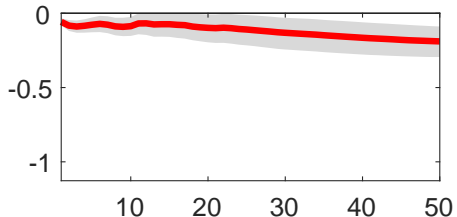
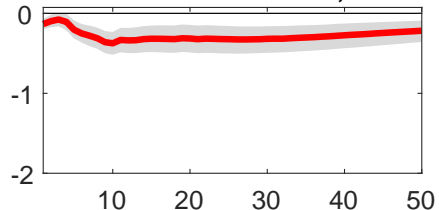




### orthogonalized MPS instrument



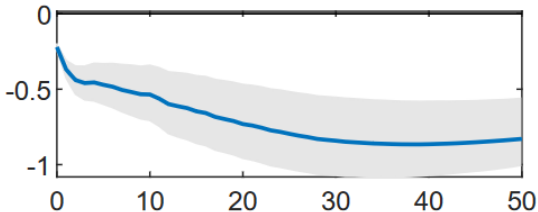
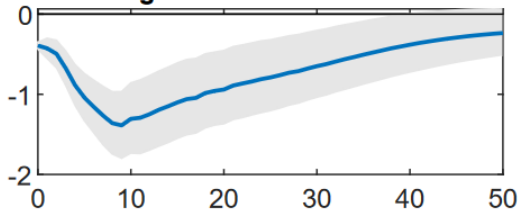
### Bauer and Swanson, NS



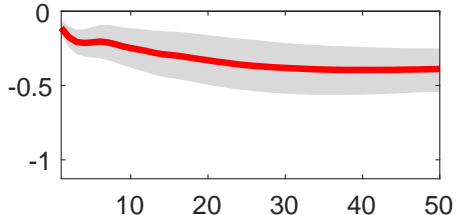
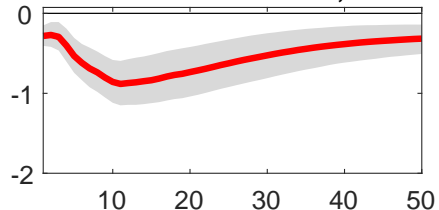




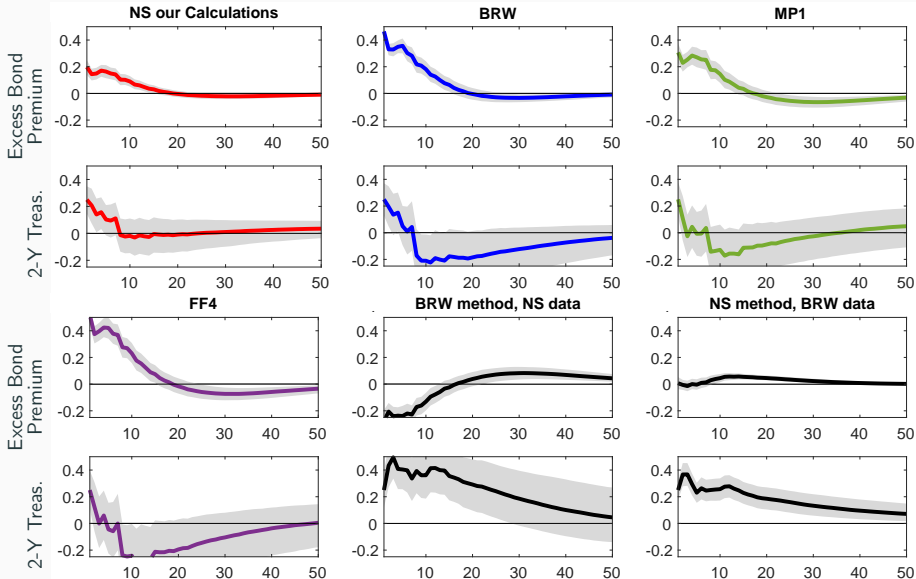
### orthogonalized MPS instrument



### Bauer and Swanson, NS



# VAR IRFs, 25 bps Monetary Shock



## Recursive vs. Full Sample

	NS Shock	NS Shock (Recursive)	BRW Shock	BRW Shock (Recursive)
Mean	-0.004	-0.002	-0.004	-0.005
Median	0.002	0.003	-0.005	-0.006
SD	0.037	0.034	0.054	0.0678

$$\text{Corr}[\text{NS Shock}, \text{NS Shock (Recursive)}] = 0.9946$$

$$\text{Corr}[\text{BRW Shock}, \text{BRW Shock (Recursive)}] = 0.9874$$